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[The jury is still out on social media as a tool for reducing food waste a response to Young et al. \(2017\).](#)

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1 Title: The jury is still out on social media as a tool for reducing food waste a response to
2 Young et al. (2017).

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7

8 Young et al. (2017) pose an important question “Can social media be a tool for reducing
9 consumers’ food waste?” Consumer food waste is thought to account for the largest
10 proportion of all food waste in developed countries (Parfitt et al. 2010). In the 28 European
11 Union countries consumer food waste constitutes between 46.7 and 63.5% of the total
12 estimated food waste of 87.6 ± 13.7 (95% CI) million tonnes (Stenmarck et al. 2016). Efforts
13 to reduce this level of waste have increasingly become important for governments and civil
14 society in the interests of environmental sustainability and cost reduction. Young et al. (2017)
15 in collaboration with a major UK supermarket company aimed to assess the influence of
16 social media (Facebook) interventions on self-reported food waste behaviour in comparison
17 to information interventions (Asda Magazine and e-newsletter) and a control group. This was
18 done in a “field” situation and not a tightly controlled experiment; hence Young et al. (2017)
19 could be an important contribution to our understanding of behaviour change in relation to
20 food waste interventions.

21

22 Young et al. (2017) report that there was no difference in the performance of the social media
23 intervention when compared to the information interventions or to the control group. They
24 suggest that all groups (interventions and controls) showed a statistically significant reduction
25 in self-reported frequency and quantity of food waste. Despite reporting the effect size (0.01),

26 Young et al. do not discuss the magnitude of the effect. Statistical significance means very
27 little in the absence of effect size (Sullivan & Feinn 2012) and a minimal (0.01) effect size
28 means that there was very little change in behaviour. With regard to the category of food
29 wasted a statistically significant decrease in salad waste is reported and Young et al. (2017)
30 suggest this is driving the pattern observed in the frequency and quantity of food wasted over
31 the three time periods. Once again the effect size was minimal (0.01) and the magnitude was
32 not discussed.

33

34 Here, the effect size represents the magnitude of the difference between the mean of a test
35 and a control group (Sullivan & Feinn 2012). It is important to note that a small effect size
36 can be meaningful (Bayliss et al. 2015). So called “t-shirt size” effects have been criticised
37 and it is essential that one relates the effect size to the data presented as an effect size of 0.4
38 (for example) could be meaningful in one study and not in another (Kline 2009). Young et al.
39 (2017) do not provide any indication of why they consider such a small effect size
40 behaviourally significant and do not provide the data behind their work (presumably due to
41 commercial confidentialities) to allow researchers to assess this independently.

42

43 The data on the frequency of waste is scaled between 1, “Never” and 5 “Most mealtimes”
44 (the intervening values are not defined in Young et al.). The mean values range between 2.36
45 and 2.63 measured on a 5 point likert scale. We used the R programme (R Core team 2016) to
46 simulate data with a similar structure to that of Young et al. (2017) using the means and
47 standard deviations as presented in Table 1 of their Appendix (all R code is available at:
48 https://osf.io/sqd8g/?view_only=27b3f2c5f1684a388ec59c0d100e7a3b). We produced 10000
49 simulated datasets (Figure 1a) and tested these with one-way repeated measures ANOVA and
50 then extracted the p values for these tests (Figure 1b). The distribution of the datasets (Figure

51 1a) shows that the different time periods greatly overlap. Only 46.6% of the 10000 tests run
52 resulted in a p value less than 0.05 (Figure 1b).

53

54 Young et al. (2017) also report the statistically significant results of t-tests comparing time
55 periods for different interventions. For example, again for the frequency of food wasted,
56 those people exposed to the Facebook intervention reported a change in behaviour from Time
57 1 (M=2.47, SD=0.91) to Time 3 (M=2.41, SD=0.91). Again using simulation (10000
58 iterations) we applied t-tests to the data. Only 17.07% of the 10000 tests resulted in a p value
59 less than 0.05 (Figure 2a). The minimal absolute difference between means to achieve a
60 statistically significant result (i.e. $p < 0.05$) is around 0.1 of a likert scale (Figure 2b).

61

62 Hence, from the data that are presented in Young et al. (2017) and our simulations we would
63 conclude that there was a small statistically significant effect but no behaviourally significant
64 effect of the interventions and of time on food waste behaviour. It is clear from our
65 simulations (Figure 1b) that the sample size ($n = 2018$) was too small to adequately identify
66 an effect if one was there. This in combination with the small effect size and the reliance on
67 self-reported measures of food waste (which is acknowledged by Young et al. 2017) increases
68 the risk of bias.

69

70 Young et al. (2017) suggest that their paper shows that “social media...cannot replicate
71 enough of the interaction shown by face to face social influence interventions to change
72 reported behaviour more than the control group (those that did not see the interventions)”.
73 This statement is premature considering the weight of knowledge that has been accumulated
74 in the behaviour change literature in fields such as psychology and medicine as over the past
75 decade. Meta-analysis has consistently reported small but positive effect sizes of online

76 interventions on behaviour change (Wantland et al. 2004; Barak et al. 2007; Maher et al.
77 2014; Short et al. 2014). The heterogeneity observed in these meta-analytical studies has been
78 attributed to the type and number of behaviour change techniques employed. Using
79 individually targeted interventions online with repeated reminders is more effective than a
80 single non-targeted approach (Short et al. 2014).

81

82 Rather than suggesting that social media cannot be used as an effective behaviour change
83 agent in the realm of food waste we suggest that Young et al. (2017) well illustrates the
84 importance of evidence-synthesis. The lack of behaviour change from a small sample of
85 people (albeit a sample size that is typical in consumer research) in a study with an untargeted
86 intervention provides one small piece of the jigsaw. The jury is still out on the potential for
87 social media to influence behaviour change and hence reduce food waste but it is imperative
88 that evidence still be collected and a variety of intervention strategies assessed. Disregarding
89 social media as a potential effective intervention on the basis of any single study would be
90 irresponsible and should not be advocated. Lack of evidence synthesis coupled with
91 (over)reliance on p values may be seriously distorting the evidence-base in this important
92 area of consumer behaviour .

93

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126 **Figure legends**

127 Figure 1. a) The distribution of data for the frequency of waste in Time period 1 to 3 for the
128 10000 simulations of data based on the means and standard deviations reported in Young et
129 al. (2017). b) The distribution of p-values for one-way repeated measures ANOVA on the
130 10000 simulated datasets. The red dashed line indicates $\alpha = 0.05$.

131

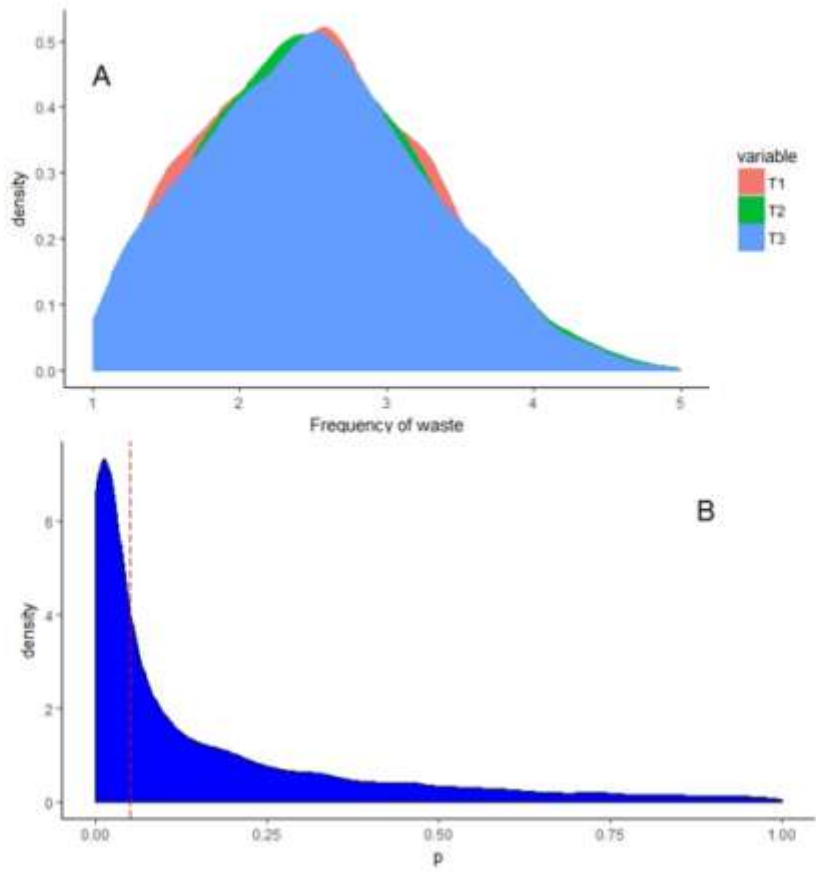
132 Figure 2. a) The distribution of p values for t-tests on 10000 simulated datasets for frequency
133 of food waste in response to the Facebook intervention. b) The absolute difference between
134 means needed to produce a statistically significant result for the 10000 simulated t-tests. The
135 red line indicates $\alpha = 0.05$.

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